1.017/1.010 Class 15 Confidence Intervals

Interval Estimates

Parameter estimates computed from a random sample $x_1, x_2, ..., x_N$ can vary around the unknown true value. For any given estimate, we seek a **two-sided confidence interval** that is likely to include the true value *a*:

$$a_L \leq a \leq a_U$$

 $[a_L, a_U]$ is called an interval estimate.

Standardized Statistics

Interval estimate is often derived from a **standardized statistic**. This is a random variable that depends on both the unknown true value *a* and its estimate \hat{a} .

An example is the **z statistic**:

$$z(\hat{a},a) = \frac{\hat{a}-a}{SD(\hat{a})}$$

If the estimate is unbiased E[z] = 0 and Var[z] = 1, for any x or \hat{a} probability distribution with finite moments (prove).

Example:

Suppose:

a = E[x] = mean of the *x* probability distribution $\hat{a} = m_x$ = sample mean (of random sample outcome $x_1, x_2, ..., x_N$).

Then:

$$z(\hat{a}, a) = \frac{\boldsymbol{m}_x - E[\boldsymbol{x}]}{SD(\boldsymbol{m}_x)} = \frac{\boldsymbol{m}_x - E[\boldsymbol{x}]}{\frac{SD(\boldsymbol{x})}{\sqrt{N}}}$$

Deriving Interval Estimates

If we know the probability distribution of the standardized statistic z we can derive an interval estimate. Specify the probability 1- α that z falls in the interval [z_L , z_U] for a **given value** of a (e.g. 0.95):

$$P[z_L \le \mathbf{z}(\hat{\mathbf{a}}, a) \le z_U] = 1 - \alpha$$

Suppose that $[z_L, z_U]$ is selected so that the probability that *z* lies above the interval is the same as the probability that it lies below the interval. This gives a **two-sided** interval $[z_L, z_U]$ for *z*:

$$P[z(\hat{a}, a) \le z_L] = F_z(z_L) = \frac{\alpha}{2}$$
$$P[z(\hat{a}, a) \ge z_U] = 1 - F_z(z_U) = \frac{\alpha}{2}$$
$$z_L = F_Z^{-1} \left[\frac{\alpha}{2}\right] \qquad z_U = F_Z^{-1} \left[1 - \frac{\alpha}{2}\right]$$

From the Central Limit Theorem we know that \hat{a} is normal and that z has a **unit normal** distribution [i.e. $z \sim N(0,1)$] for **large sample sizes**. In this case the CDF $F_z(z)$ and its inverse can be evaluated from standard normal distribution tables or with the MATLAB functions normcdf and norminv.

If
$$1-\alpha = 0.95$$
 then $z_L = -1.96$ and $z_U = +1.96$.

Substitute the definition of *z* to obtain the corresponding two-sided interval for \hat{a} :

$$P[z_{L} \leq \frac{\hat{a} - a}{SD[\hat{a}]} \leq z_{U}] = P[a + z_{L}SD[\hat{a}] \leq \hat{a} \leq a + z_{U}SD[\hat{a}]] = 1 - \alpha$$
$$\hat{a} \geq a + z_{L}SD[\hat{a}] \qquad \hat{a} \leq a + z_{U}SD[\hat{a}]$$

Probability is 1- α that actual sample estimate \hat{a} lies in this interval.

Now suppose that this relatively likely event occurs when the outcome of a particular experiment (i.e. the \hat{a} obtained from a particular random sample) is \hat{a} . Then the true *a* must satisfy the following inequality:

$$\hat{a} - z_U SD[\hat{a}] \le a \le \hat{a} - z_L SD[\hat{a}]$$

This gives the desired 1- α confidence interval for *a*:

 $a_L = \hat{a} - z_U SD[\hat{a}]$ $a_U = \hat{a} - z_L SD[\hat{a}]$

We can obtain $SD[\hat{a}]$ in two ways:

1. Derive directly from the **definition of the estimator** $\hat{a} = \hat{a}(x_1, x_2, ..., x_N)$ (not always possible). This usually requires replacing population statistics [e.g. *Std*(*x*)] by sample statistics [e.g. *s_x*]

2. **Stochastic simulation**, using $a = \hat{a}$ in random number generator (usually possible but not exact). This generally requires an assumption about the form of the underlying distribution $F_x(x)$.

The confidence interval is wider for larger $SD[\hat{a}]$

The confidence interval is wider for larger $1-\alpha$ (e.g. 99%)

Summary

To derive a two-sided confidence interval:

- 1. Specify significance level α
- 2. Compute estimate \hat{a} from the data
- 3. Compute *Std* [\hat{a}] in one of two ways:
 - If possible, relate $Std(\hat{a})$ to Std(x) and use the approximation $Std(x) \approx s_x$ (i.e assume the unknown population standard deviation is equal to the sample standard deviation computed from data).
 - Otherwise, derive $Std(\hat{a})$ using stochastic simulation
- 4. Compute z_L and z_U from and the specified α , assuming an appropriate form for the CDF $F_x(x)$
- 5. Apply the two-sided confidence interval formula

Example – Large-sample two-sided confidence interval for the population mean:

Consider the sample mean m_x , used to estimate the population mean E[x]. In this case, a = E[x] and $\hat{a} = m_x$

Use result from Class 13 to derive $SD[m_x]$ directly, replacing SD[x] by the sample standard deviation s_x :

$$SD[\boldsymbol{m}_{x}] = \frac{SD[\boldsymbol{x}]}{\sqrt{N}} \approx \frac{s_{x}}{\sqrt{N}}$$

So the large sample (assume z is normal) two-sided 95% confidence interval for the population mean is:

$$m_x - 1.96 \frac{s_x}{\sqrt{N}} \le E[\mathbf{x}] \le m_x + 1.96 \frac{s_x}{\sqrt{N}}$$

Suppose:

 $[\mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_{10}] = [0.1 \ 2.9 \ 1.0 \ 1.4 \ 0.23 \ 0.54 \ 1.57 \ 8.0 \ 0.40 \ 1.6]$

Then $m_x = 1.77$, $s_x = 2.34$, N = 10 and:

$$1.77 - 1.96 \frac{2.34}{\sqrt{10}} \le E[\mathbf{x}] \le 1.77 + 1.96 \frac{2.34}{\sqrt{10}}$$
$$0.32 \le E[\mathbf{x}] \le 3.23$$

Suppose we don't know the probability distribution of z (e.g. because the sample size is too small to justify using the Central Limit Theorem) but we know the distribution of x (except for a few unknown parameters).

Then we can approximate $F_z(z)$ with a virtual experiment, replacing unknown parameters with estimates computed from the random sample. Once this is done we can plot $F_z(z)$ vs z and identify from the plot the F_z^{-1} values needed to derive z_L and z_U . Usually we make a large sample (normal) assumption.

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