# Implementation of IV and Two-Stage Least Squares 14.32, Spring 2007 

## Review of IV

In lecture we saw the schooling model

$$
y_{i}=\alpha+\rho S_{i}+\gamma A_{i}+\varepsilon_{i}
$$

where $y_{i}$ is log income of individual $i, S_{i}$ is schooling, and $A_{i}$ is unmeasured ability. Since ability is unobservable, the OLS regression becomes

$$
y_{i}=\alpha+\rho S_{i}+u_{i}
$$

where $u_{i}=\gamma A_{i}+\varepsilon_{i}$. From the omitted variables bias formula, the coefficient we estimate has $p \lim (\hat{\rho})=\rho+\gamma \beta_{A S}$, where $\beta_{A S}$ is the coefficient of the bivarite regression of ability on schooling.

One way to deal with this problem is to put in enough covariates so that schooling is no longer correlated with the unobserved error term. This works in some cases, but often you can't find the appropriate covariates, e.g., how can you measure ability? Some people have used test scores, but these results are not 100\% convincing.

Instrumental variables estimation gets around the omitted variables bias problem by finding another variable, called an instrument, which affects the variable of interest but nothing else in the regression. More formally, you find some instrument $z$ where $\operatorname{cov}\left(z_{i}, u_{i}\right)=0$. Then, you can back out the parameter of interest using covariances:

$$
\begin{aligned}
\operatorname{cov}\left(y_{i}, z_{i}\right) & =\operatorname{cov}\left(\alpha+\rho S_{i}+u_{i}, z_{i}\right)=\rho \operatorname{cov}\left(S_{i}, z_{i}\right) \\
& \Rightarrow \frac{\operatorname{cov}\left(y_{i}, z_{i}\right)}{\operatorname{cov}\left(S_{i}, z_{i}\right)}=\frac{\beta_{y z}}{\beta_{S z}}=\rho
\end{aligned}
$$

From the law of large numbers, the sample analog will give you a consistent estimate, i.e.,

$$
p \lim \frac{\widehat{\operatorname{cov}\left(y_{i}, z_{i}\right)}}{\widehat{\operatorname{cov}\left(S_{i}, z_{i}\right)}}=p \lim \frac{\widehat{\beta_{y z}}}{\widehat{\beta_{S z}}}=\rho
$$

The important assumptions for the instrument $z$ are $\operatorname{cov}\left(z_{i}, S_{i}\right) \neq 0$ and $\operatorname{cov}\left(z_{i}, u_{i}\right)=0$. In words we can state these two assumptions as:

1. The insturment $z$ is correlated with the endogenous variable $S$ (there is a first stage).
2. The instrument $z$ only affects $y$ through the variable $S$ (exclusion restriction).

## Two-Stage Least Squares

Practically, IV is usually implemented through a process called two-stage least squares (2SLS). This is actually a simple type of simultaneous equations problem. Let's set up the relationships between $y, S, z$ and $u$ as

$$
\begin{aligned}
& y_{i}=\alpha+\rho S_{i}+u_{i} \\
& S_{i}=\kappa+\delta z_{i}+v_{i}
\end{aligned}
$$

To run 2SLS:

1. Regress $S$ on $z$
2. Using the estimated parameters, estimate $\hat{S}$
3. Regress $y$ on $\hat{S}$
4. The estimated $\hat{\rho}$ is your IV estimator.

It turns out that through this process the estimated $\hat{\rho}$ is numerically equivalent to the IV estimator developed above, $\frac{\widehat{\operatorname{cov}(y, z)}}{\widehat{\operatorname{cov}(z, S)}}$. Here's the proof (supressing some hats and $i^{\prime} s$ to make it easier to read):

$$
\begin{aligned}
\hat{\rho}_{2 S L S} & =\frac{\operatorname{cov}(y, \hat{S})}{\operatorname{var}(\hat{S})} \\
& =\frac{\operatorname{cov}\left(y, \frac{\operatorname{cov}(S, z)}{\operatorname{var}(z)} z+\hat{\kappa}\right)}{\operatorname{var}\left(\frac{\operatorname{cov}(S, z)}{\operatorname{var}(z)} z+\hat{\kappa}\right)} \\
& =\frac{\frac{\operatorname{cov}(S, z)}{\operatorname{var}(z)} \operatorname{cov}(y, z)}{\left(\frac{\operatorname{cov}(S, z)}{\operatorname{var}(z)}\right)^{2} \operatorname{var}(z)} \\
& =\frac{\operatorname{cov}(y, z)}{\operatorname{cov}(S, z)}=\hat{\rho}_{I V}
\end{aligned}
$$

## Computing 2SLS Estimates

2SLS makes computing IV estimators easy. Here's an example with SAS. As we have seen in class, there are a number of instruments one can use to estimate the wage-schooling equation

$$
y_{i}=\alpha+\rho S_{i}+u_{i}
$$

An alternative (and clever) instrument for schooling was developed by Card (1995). His insight was that people who live closer to colleges have a lower cost of attending school, and therefore higher education levels. For his instrument he constructed an indicator for whether the individual lived near a four-year college. The quality of this instrument has been the subject of debate, and you can make your own judgement as to its own validity.
To compute 2SLS, we first estimate the first-stage regression:

$$
S_{i}=\kappa+\delta * \text { nearc }_{i}+v_{i}
$$

Then take the fitted values and estimate

$$
y_{i}=\alpha+\rho \hat{S}_{i}+\zeta_{i}
$$

Note that the standard errors from the second OLS estimation need to be corrected for 2 SLS. The basic idea of this is that once we have a consistent estimate of $\hat{\rho}$, we use $S$, not $\hat{S}$, to compute the residuals. SAS does this automatically with PROC SYSLIN.

```
/*2SLS DEMONSTRATION*/
data one;
    set 'card';
/*OLS REGRESSION*/
proc reg data=one;
    model lwage=educ;
    title "Simple OLS";
/*FIRST STAGE OF 2SLS REGRESSION*/
proc reg data=one;
    model educ=nearc4;
    output out=one
    p=educhat;
    title "First Stage" ;
/*SECOND STAGE*/
proc reg data=one;
    model lwage=educhat;
    title "Second Stage";
/*DOING IT ALL AT ONCE USING PROC SYSLIN*/
proc syslin data=one 2sls;
    endogenous educ;
    instruments nearc4;
    model lwage=educ;
    title "2SLS with PROC SYSLIN";
run;
```

Cite as: James Berry, course materials for 14.32 Econometrics, Spring 2007. MIT OpenCourseWare (http://ocw.mit.edu/), Massachusetts Institute of Technology. Downloaded on [DD Month YYYY].


Cite as: James Berry, course materials for 14.32 Econometrics, Spring 2007. MIT OpenCourseWare (http://ocw.mit.edu/), Massachusetts Institute of Technology. Downloaded on [DD Month YYYY].


Cite as: James Berry, course materials for 14.32 Econometrics, Spring 2007. MIT OpenCourseWare (http://ocw.mit.edu/), Massachusetts Institute of Technology. Downloaded on [DD Month YYYY].

# The REG Procedure 

Model: MODEL1
Dependent Variable: lwage
Number of Observations Read 3010
Number of Observations Used 3010

| Analysis of Variance |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Source | DF | Sum of Squares | Mean Square | F Value | Pr $>\mathrm{F}$ |
| Model | 1 | 15.86603 | 15.86603 | 82.74 | $<.0001$ |
| Error | 3008 | 576.77560 | 0.19175 |  |  |
| Corrected Total | 3009 | 592.64163 |  |  |  |


| Root MSE | 0.43789 | R-Square | 0.0268 |
| :--- | :--- | :--- | :--- |
| Dependent Mean | 6.26183 | Adj R-Sq | 0.0264 |
| Coeff Var | 6.99299 |  |  |


| Variable | Label | DF | Parameter <br> Estimate | Standard <br> Error |
| :--- | :--- | :---: | ---: | :--- |
| Intercept | Intercept |  |  | Value |

Cite as: James Berry, course materials for 14.32 Econometrics, Spring 2007. MIT OpenCourseWare (http://ocw.mit.edu/), Massachusetts Institute of Technology. Downloaded on [DD Month YYYY].

The SYSLIN Procedure
Two-Stage Least Squares Estimation

| Model | lwage |
| :--- | :--- |
| Dependent Variable | lwage |

Analysis of Variance


Cite as: James Berry, course materials for 14.32 Econometrics, Spring 2007. MIT OpenCourseWare (http://ocw.mit.edu/), Massachusetts Institute of Technology. Downloaded on [DD Month YYYY].

