Sums of Random Variables

Consider *n* RVs x_i and let $s \equiv \sum_{i=1}^n x_i$.

If the RVs are statistically independent, then

$$\langle s \rangle = \sum_{i} \langle x_i \rangle$$

$$Var(s) = \sum_{i} Var(x_i)$$

• The individual $p(x_i)$ could be quite different

• Both continuous and discrete RVs could be present

 \bullet True for any n

• Even if one RV dominates the sum

Results have a special meaning when

1) The means are finite $(\neq 0)$

2) The variances are finite $(\neq \infty)$

3) No subset dominates the sum

4) n is large



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Given p(x, y), find $p(s \equiv x + y)$



B
$$P_s(\alpha) = \int_{-\infty}^{\infty} d\zeta \int_{-\infty}^{\alpha-\zeta} d\eta \ p_{x,y}(\zeta,\eta)$$

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C
$$p_s(\alpha) = \int_{-\infty}^{\infty} d\zeta \ p_{x,y}(\zeta, \alpha - \zeta)$$

This is a general result; x and y need not be S.I.

Application to the Jointly Gaussian RVs in Section 2 shows that p(s) is a Gaussian with zero mean and a Variance = $2\sigma^2(1 + \rho)$.

In the special case that x and y are S.I.

$$p_s(\alpha) = \int_{-\infty}^{\infty} d\zeta \ p_x(\zeta) \ p_y(\alpha - \zeta) = \int_{-\infty}^{\infty} d\zeta' \ p_x(\alpha - \zeta') \ p_y(\zeta')$$

The mathematical operation is called "convolution".

$$p\otimes q\equiv \int_{-\infty}^{\infty}p(z)q(x-z)dz=f(x).$$

Example

Given:

$$p(z) = \frac{1}{n!a} (z/a)^n \exp(-z/a)$$

$$q(z) = \frac{1}{m!a} (z/a)^m \exp(-z/a)$$



 $0 < z \text{ and } n, m = 0, 1, 2, \cdots$

Find: $p \otimes q$

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$$p \otimes q = \frac{1}{n!m!} \frac{1}{a^2} \int_0^x \left(\frac{z}{a}\right)^n \left(\frac{x-z}{a}\right)^m e^{-z/a} e^{-(x-z)/a} dz$$

$$= \frac{1}{n!m!} \frac{1}{a} \left(\frac{1}{a}\right)^{n+m+1} e^{-x/a} \int_0^x z^n (x-z)^m dz$$

$$= \frac{1}{n!m!} \frac{1}{a} \left(\frac{x}{a}\right)^{n+m+1} e^{-x/a} \underbrace{\int_{0}^{1} \zeta^{n} (1-\zeta)^{m} d\zeta}_{\frac{n!m!}{(n+m+1)!}}$$

$$p \otimes q = \frac{1}{(n+m+1)!} \frac{1}{a} \left(\frac{x}{a}\right)^{n+m+1} e^{-x/a}$$

a function of the same class

Example Atomic Hydrogen Maser



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$$t_{\text{wall}}$$
 (given n stays) = $\sum_{i=1}^{n} t_i$

 $t_i \equiv$ duration of i_{th} stay on wall. Each stay is S.I.

$$p(t | 1) = (1/\tau) e^{-t/\tau}$$

$$p(t | 2) = p(t | 1) \otimes p(t | 1) = (1/\tau)(t/\tau) e^{-t/\tau}$$

 $p(t | 3) = p(t | 2) \otimes p(t | 1) = (1/2)(1/\tau)(t/\tau)^2 e^{-t/\tau}$

$$p(t | n) = \frac{1}{(n-1)!} \frac{1}{\tau} \left(\frac{t}{\tau}\right)^{n-1} e^{-t/\tau}$$



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Facts about sums of RVs

• Exact expressions for $\langle s \rangle$ and Var(s) if S.I.

•
$$p(s) = p(x) \otimes p(y)$$
 if S.I.

• p(s) slightly more complicated if not S.I.

 $\bullet \otimes$ usually changes functional form

• But not always

• Fourier techniques are very useful

Very important special case: Central Limit Theorem

- RVs are S.I.
- All have identical densities $p(x_i)$
- Var(x) is finite but $\langle x \rangle$ could be zero
- n is large



If x is continuous

$$p(s) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(s-\langle s \rangle)^2/2\sigma^2}$$
$$\langle s \rangle = n \langle x \rangle$$
$$\sigma^2 = n \sigma_x^2$$

If x is discrete in equal steps of Δx





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Non-rigorous extensions of the Central Limit Theorem

- The Gaussian can be a good practical approximation for modest values of *n*.
- The Central Limit Theorem may work even if the individual members of the sum are not identically distributed.
- The requirement that the variables be statistically independent may even be waived in some cases, particularly when *n* is very large

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